

On the Growth Rate of the Weight Distribution of Irregular Doubly-Generalized LDPC Codes

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Abstract—In this paper, an expression for the asymptotic growth rate of the number of small linear-weight codewords of irregular doubly-generalized LDPC (D-GLDPC) codes is derived. The expression is compact and generalizes existing results for LDPC and generalized LDPC (GLDPC) codes. Assuming that there exist check and variable nodes with minimum distance 2, it is shown that the growth rate depends only on these nodes. An important connection between this new result and the stability condition of D-GLDPC codes over the BEC is highlighted. Such a connection, previously observed for LDPC and GLDPC codes, is now extended to the case of D-GLDPC codes.

I. INTRODUCTION

Recently, low-density parity-check (LDPC) codes have been intensively studied due to their near-Shannon-limit performance under iterative belief-propagation decoding. Binary regular LDPC codes were first proposed by Gallager in 1963 [1]. In the last decade the capability of irregular LDPC codes to outperform regular ones in the waterfall region of the performance curve and to asymptotically approach (or even achieve) the communication channel capacity has been recognized and deeply investigated (see for instance [2], [3], [4], [5], [6], [7]).

It is usual to represent an LDPC code as a bipartite graph, i.e., as a graph where the nodes are grouped into two disjoint sets, namely, the variable nodes (VNs) and the check nodes (CNs), such that each edge may only connect a VN to a CN. The bipartite graph is also known as a Tanner graph [8]. In the Tanner graph of an LDPC code, a generic degree- q VN can be interpreted as a length- q repetition code, as it repeats q times its single information bit towards the CNs. Similarly, a degree- s CN of an LDPC code can be interpreted as a length- s single parity-check (SPC) code, as it checks the parity of the s VNs connected to it.

The growth rate of the weight distribution of Gallager's regular LDPC codes was investigated in [1]. The analysis demonstrated that, provided that the smallest VN degree is at least 3, for large enough codeword length N , the expected minimum distance of a randomly chosen code in the ensemble is a linear function of N .

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More recently, the study of the weight distribution of binary LDPC codes has been extended to irregular ensembles. Important works in this area are [9], [10], [11]. In [11] a complete solution for the growth rate of the weight distribution of binary irregular LDPC codes was developed. One of the main results of [11] is a connection between the expected behavior of the weight distribution of a code randomly chosen from the ensemble and the parameter $\lambda'(0)\rho'(1)$, $\lambda(x)$ and $\rho(x)$ being the edge-perspective VN and CN degree distributions, respectively. More specifically, it was shown that for a code randomly chosen from the ensemble, one can expect an exponentially small number of small linear-weight codewords if $0 \leq \lambda'(0)\rho'(1) < 1$, and an exponentially large number of small linear-weight codewords if $\lambda'(0)\rho'(1) > 1$.

This result establishes a connection between the statistical properties of the weight distribution of binary irregular LDPC codes and the stability condition of binary irregular LDPC codes over the binary erasure channel (BEC) [3], [4]. If q^* denotes the LDPC asymptotic iterative decoding threshold over the BEC, the stability condition states that we always have

$$q^* \leq [\lambda'(0)\rho'(1)]^{-1}. \quad (1)$$

Prior to the rediscovery of LDPC codes, binary generalized LDPC (GLDPC) codes were introduced by Tanner in 1981 [8]. A GLDPC code generalizes the concept of an LDPC code in that a degree- s CN may in principle be any (s, h) linear block code, s being the code length and h the code dimension. Such a CN accounts for $s-h$ linearly independent parity-check equations. A CN associated with a linear block code which is not a SPC code is said to be a *generalized CN*. In [8] *regular* GLDPC codes (also known as Tanner codes) were investigated, these being GLDPC codes where the VNs are all repetition codes of the same length and the CNs are all linear block codes of the same type.

The growth rate of the weight distribution of binary GLDPC codes was investigated in [12], [13], [14], [15]. In [12] the growth rate is calculated for Tanner codes with BCH check component codes and length-2 repetition VNs, leading to an asymptotic lower bound on the minimum distance. The same lower bound is developed in [13] assuming Hamming CNs and length-2 repetition VNs. Both works extend the approach developed by Gallager in [1, Chapter 2] to show that, for sufficiently large N , the minimum distance is a linear function of N . The growth rate of the number of small weight codewords for GLDPC codes with a uniform CN set (all CN of the same type) and an irregular VN

set (repetition VNs with different lengths) is investigated in [14]. It is shown that for sufficiently large N , a minimum distance increasing linearly with N is expected when either the uniform CN set is composed of linear block codes with minimum distance at least 3, or the minimum length of the repetition VNs is 3. On the other hand, if the minimum distance of the CNs and the minimum length of the repetition VNs are both equal to 2, then for a randomly selected GLDPC code in the ensemble we expect a minimum distance growing as a linear or sublinear function of N (for large N) depending on the sign of the first order coefficient in the growth rate Taylor series expansion. The results developed in [14] were further extended in [15] to GLDPC ensembles with an irregular CN set (CNs of different types). It was there proved that, provided that there exist CNs with minimum distance 2, a parameter $\lambda'(0)C$, generalizing the parameter $\lambda'(0)\rho'(1)$ of LDPC code ensembles, plays in the context of the weight distribution of GLDPC codes the same role played by $\lambda'(0)\rho'(1)$ in the context of the weight distribution of LDPC codes. The parameter C is defined in Section III.

Interestingly, this latter results extends to binary GLDPC codes the same connection between the statistical properties of the weight distribution of irregular codes and the stability condition over the BEC. In fact, it was shown in [16] that the stability condition of binary irregular GLDPC codes over the BEC is given by

$$q^* \leq [\lambda'(0)C]^{-1}. \quad (2)$$

Generalized LDPC codes represent a promising solution for low-rate channel coding schemes, due to an overall rate loss introduced by the generalized CNs [17]. Doubly-generalized LDPC (D-GLDPC) codes generalize the concept of GLDPC codes while facilitating much greater design flexibility in terms of code rate [18] (an analogous idea may be found in the previous work [19]). In a D-GLDPC code, the VNs as well as the CNs may be of any generic linear block code types. A degree- q VN may in principle be any (q, k) linear block code, q being the code length and k the code dimension. Such a VN is associated with k D-GLDPC code bits. It interprets these bits as its local information bits and interfaces to the CN set through its q local code bits. A VN which corresponds to a linear block code which is not a repetition code is said to be a *generalized VN*. A D-GLDPC code is said to be *regular* if all of its VNs are of the same type and all of its CNs are of the same type and is said to be *irregular* otherwise.

In this paper the growth rate of the weight distribution of binary irregular D-GLDPC codes is analyzed for small weight codewords. It is shown that, provided there exist both VNs and CNs with minimum distance 2, a parameter $1/P^{-1}(1/C)$ discriminates between an asymptotically small and an asymptotically large expected number of small linear-weight codewords in a D-GLDPC code randomly drawn from a given irregular ensemble (the function $P(x)$ is defined in Section III). The parameter $1/P^{-1}(1/C)$ generalizes the above mentioned parameters $\lambda'(0)\rho'(1)$ and $\lambda'(0)C$ to the case where both generalized VNs and generalized CNs are

present. The obtained result also represents the extension to the D-GLDPC case of the previously recalled connection with the stability condition over the BEC. In fact, it was proved in [16] that the stability condition of D-GLDPC codes over the BEC is given by

$$q^* \leq P^{-1}(1/C). \quad (3)$$

The paper is organized as follows. Section II defines the D-GLDPC ensemble of interest, and introduces some definitions and notation pertaining to this ensemble. Section III defines further terms regarding the VNs and CNs which compose the D-GLDPC codes in the ensemble. Finally, Section IV states and proves the main result of the paper regarding the growth rate of the weight distribution.

II. IRREGULAR DOUBLY-GENERALIZED LDPC CODE ENSEMBLE

We define a D-GLDPC code ensemble \mathcal{M}_n as follows, where n denotes the number of VNs. There are n_c different CN types $t \in I_c = \{1, 2, \dots, n_c\}$, and n_v different VN types $t \in I_v = \{1, 2, \dots, n_v\}$. For each CN type $t \in I_c$, we denote by h_t , s_t and r_t the CN dimension, length and minimum distance, respectively. For each VN type $t \in I_v$, we denote by k_t , q_t and p_t the VN dimension, length and minimum distance, respectively. For $t \in I_c$, ρ_t denotes the fraction of edges connected to CNs of type t . Similarly, for $t \in I_v$, λ_t denotes the fraction of edges connected to VNs of type t . Note that all of these variables are independent of n .

The polynomials $\rho(x)$ and $\lambda(x)$ are defined by

$$\rho(x) = \sum_{t \in I_c} \rho_t x^{s_t - 1}$$

and

$$\lambda(x) = \sum_{t \in I_v} \lambda_t x^{q_t - 1}.$$

If E denotes the number of edges in the Tanner graph, the number of CNs of type $t \in I_c$ is then given by $E\rho_t/s_t$, and the number of VNs of type $t \in I_v$ is then given by $E\lambda_t/q_t$. Denoting as usual $\int_0^1 \rho(x) dx$ and $\int_0^1 \lambda(x) dx$ by $\int \rho$ and $\int \lambda$ respectively, we see that the number of edges in the Tanner graph is given by

$$E = \frac{n}{\int \lambda}$$

and the number of CNs is given by $m = E \int \rho$. Therefore, the fraction of CNs of type $t \in I_c$ is given by

$$\gamma_t = \frac{\rho_t}{s_t \int \rho} \quad (4)$$

and the fraction of VNs of type $t \in I_v$ is given by

$$\delta_t = \frac{\lambda_t}{q_t \int \lambda} \quad (5)$$

Also the length of any D-GLDPC codeword in the ensemble is given by

$$N = \sum_{t \in I_v} \left(\frac{E\lambda_t}{q_t} \right) k_t = \frac{n}{\int \lambda} \sum_{t \in I_v} \frac{\lambda_t k_t}{q_t}. \quad (6)$$

Note that this is a linear function of n . Similarly, the total number of parity-check equations for any D-GLDPC code in the ensemble is given by

$$M = \frac{m}{\int \rho} \sum_{t \in I_c} \frac{\rho_t (s_t - h_t)}{s_t}.$$

A member of the ensemble then corresponds to a permutation on the E edges connecting CNs to VNs.

The growth rate of the weight distribution of the irregular D-GLDPC ensemble sequence $\{\mathcal{M}_n\}$ is defined by

$$G(\alpha) = \lim_{n \rightarrow \infty} \frac{1}{n} \log \mathbb{E}_{\mathcal{M}_n} [N_{\alpha n}] \quad (7)$$

where $\mathbb{E}_{\mathcal{M}_n}$ denotes the expectation operator over the ensemble \mathcal{M}_n , and N_w denotes the number of codewords of weight w of a randomly chosen D-GLDPC code in the ensemble \mathcal{M}_n . The limit in (7) assumes the inclusion of only those positive integers n for which $\alpha n \in \mathbb{Z}$ and $\mathbb{E}_{\mathcal{M}_n} [N_{\alpha n}]$ is positive (i.e., where the expression whose limit we seek is well defined). Note that the argument of the growth rate function $G(\alpha)$ is equal to the ratio of D-GLDPC codeword length to the number of VNs; by (6), this captures the behaviour of codewords linear in the block length, as in [11] for the LDPC case.

Definition 2.1: An *assignment* is a subset of the edges of the Tanner graph. An assignment is said to have *weight* k if it has k elements. An assignment is said to be *check-valid* if the following condition holds: supposing that each edge of the assignment carries a 1 and each of the other edges carries a 0, each CN recognizes a valid codeword.

Definition 2.2: A *split assignment* is an assignment, together with a subset of the D-GLDPC code bits (called a *codeword assignment*). A split assignment is said to have *split weight* (u, v) if its assignment has weight v and its codeword assignment has u elements. A split assignment is said to be *check-valid* if its assignment is check-valid. A split assignment is said to be *variable-valid* if the following condition holds: supposing that each edge of its assignment carries a 1 and each of the other edges carries a 0, and supposing that each D-GLDPC code bit in the codeword assignment is set to 1 and each of the other code bits is set to 0, each VN recognizes an input word and the corresponding valid codeword.

Note that for any D-GLDPC code, there is a bijective correspondence between the set of D-GLDPC codewords and the set of split assignments which are both variable-valid and check-valid.

III. FURTHER DEFINITIONS AND NOTATION

The weight enumerating polynomial for CN type $t \in I_c$ is given by

$$\begin{aligned} A^{(t)}(x) &= \sum_{u=0}^{s_t} A_u^{(t)} x^u \\ &= 1 + \sum_{u=r_t}^{s_t} A_u^{(t)} x^u. \end{aligned}$$

Here $A_u^{(t)} \geq 0$ denotes the number of weight- u codewords for CNs of type t . Note that $A_{r_t}^{(t)} > 0$ for all $t \in I_c$. The bivariate weight enumerating polynomial for VN type $t \in I_v$ is given by

$$\begin{aligned} B^{(t)}(x, y) &= \sum_{u=0}^{k_t} \sum_{v=0}^{q_t} B_{u,v}^{(t)} x^u y^v \\ &= 1 + \sum_{u=1}^{k_t} \sum_{v=p_t}^{q_t} B_{u,v}^{(t)} x^u y^v. \end{aligned}$$

Here $B_{u,v}^{(t)} \geq 0$ denotes the number of weight- v codewords generated by input words of weight u , for VNs of type t . Also, for each $t \in I_v$, corresponding to the polynomial $B^{(t)}(x, y)$ we denote the sets

$$S_t = \{(i, j) \in \mathbb{Z}^2 : B_{i,j}^{(t)} > 0\} \quad (8)$$

and

$$S_t^- = S_t \setminus \{(0, 0)\}. \quad (9)$$

We denote the smallest minimum distance over all CN types by

$$r = \min\{r_t : t \in I_c\}$$

and the set of CN types with this minimum distance by

$$X_c = \{t \in I_c : r_t = r\}.$$

We also define

$$C_t = \frac{r_t A_{r_t}^{(t)}}{s_t}$$

for each $t \in I_c$, and

$$C = \sum_{t \in X_c} \rho_t C_t. \quad (10)$$

We note that $C_t > 0$ for all $t \in I_c$, so $C > 0$.

Similarly, we denote the smallest minimum distance over all VN types by

$$p = \min\{p_t : t \in I_v\}$$

and the set of VN types with this minimum distance by

$$X_v = \{t \in I_v : p_t = p\}.$$

In the specific case where $p = 2$, we also introduce the following definitions. For each $t \in X_v$, define the set $L_t = \{i \in \mathbb{Z} : B_{i,2}^{(t)} > 0\}$ – note that these sets are nonempty. Also define the polynomial $P(x)$ by

$$P(x) = \sum_{t \in X_v} \lambda_t \sum_{i \in L_t} \frac{2B_{i,2}^{(t)}}{q_t} x^i. \quad (11)$$

Since all the coefficients of $P(x)$ are positive, $P(x)$ is monotonically increasing on $[0, \infty)$ and therefore its inverse $P^{-1}(x)$ is well-defined and unique on this interval. Note that in the case $r = p = 2$, both C and the polynomial $P(x)$ depend only on the CNs and VNs with minimum distance equal to 2.

Finally, note that throughout this paper, the notation $e = \exp(1)$ denotes Napier's number.

IV. GROWTH RATE FOR DOUBLY-GENERALIZED LDPC CODE ENSEMBLE

The following theorem constitutes the main result of the paper.

Theorem 4.1: Consider an irregular D-GLDPC code ensemble sequence \mathcal{M}_n satisfying $r = p = 2$. For sufficiently small α , the growth rate of the weight distribution is given by

$$G(\alpha) = \alpha \log \left[\frac{1}{P^{-1}(1/C)} \right] + O(\alpha^2). \quad (12)$$

The theorem is proved next. For ease of presentation, the proof is broken into four parts.

A. Number of check-valid assignments of weight ϵm over γm CNs of type $t \in I_c$

Consider γm CNs of the same type $t \in I_c$. Using generating functions¹, the number of check-valid assignments (over these CNs) of weight ϵm is given by

$$N_{c,t}^{(\gamma m)}(\epsilon m) = \text{Coeff} \left[\left(A^{(t)}(x) \right)^{\gamma m}, x^{\epsilon m} \right]$$

where $\text{Coeff} [p(x), x^c]$ denotes the coefficient of x^c in the polynomial $p(x)$. We now use the following result, which appears as Lemma 19 in [11]:

Lemma 4.2: Let $A(x) = 1 + \sum_{u=c}^d A_u x^u$, where $1 \leq c \leq d$, be a polynomial satisfying $A_c > 0$ and $A_u \geq 0$ for all $c < u \leq d$. Then, for sufficiently small ξ ,

$$\lim_{\ell \rightarrow \infty} \frac{1}{\ell} \log \text{Coeff} \left[(A(x))^\ell, x^{\xi \ell} \right] = \frac{\xi}{c} \log \left(\frac{ecA_c}{\xi} \right) + O(\xi^2) \quad (13)$$

The limit in (13) assumes the inclusion of only those positive integers ℓ for which $\xi \ell \in \mathbb{Z}$ and $\text{Coeff} [(A(x))^\ell, x^{\xi \ell}]$ is positive (i.e., where the expression whose limit we seek is well defined).

Applying this lemma by substituting $A(x) = A^{(t)}(x)$, $\ell = \gamma m$ and $\xi = \epsilon/\gamma$, we obtain that with γ fixed, as $m \rightarrow \infty$ we have, for sufficiently small ϵ ,

$$N_{c,t}^{(\gamma m)}(\epsilon m) = \text{Coeff} \left[\left(A^{(t)}(x) \right)^{\gamma m}, x^{\epsilon m} \right] \rightarrow \exp \left\{ m \left[\frac{\epsilon}{r_t} \log \left(\frac{er_t A_{r_t}^{(t)} \gamma}{\epsilon} \right) + O(\epsilon^2) \right] \right\} \quad (14)$$

¹Here we make use of the following general result [20]. Let a_i be the number of ways of obtaining an outcome $i \in \mathbb{Z}$ in experiment \mathcal{A} , and let b_j be the number of ways of obtaining an outcome $j \in \mathbb{Z}$ in experiment \mathcal{B} . Also let c_k be the number of ways of obtaining an outcome (i, j) in the combined experiment $(\mathcal{A}, \mathcal{B})$ with sum $i + j = k$. Then the generating functions $A(x) = \sum_i a_i x^i$, $B(x) = \sum_j b_j x^j$ and $C(x) = \sum_k c_k x^k$ are related by $C(x) = A(x)B(x)$.

B. Number of check-valid assignments of weight δm

Next we derive an expression, valid asymptotically, for the number of check-valid assignments of weight δm . For each $t \in I_c$, let $\epsilon_t m$ denote the portion of the total weight δm apportioned to CNs of type t . Then $\epsilon_t \geq 0$ for each $t \in I_c$, and $\sum_{t \in I_c} \epsilon_t = \delta$. Also denote $\epsilon = (\epsilon_1 \ \epsilon_2 \ \dots \ \epsilon_{n_c})$. The number of check-valid assignments of weight δm satisfying the constraint ϵ is obtained by multiplying the numbers of check-valid assignments of weight $\epsilon_t m$ over $\gamma_t m$ CNs of type t , for each $t \in I_c$,

$$N_c^{(\epsilon)}(\delta m) = \prod_{t \in I_c} N_{c,t}^{(\gamma_t m)}(\epsilon_t m)$$

where the fraction γ_t of CNs of type $t \in I_c$ is given by (4).

As $n \rightarrow \infty$, we have $m \rightarrow \infty$ and so we obtain using (14) that for sufficiently small δ ,

$$N_c^{(\epsilon)}(\delta m) \rightarrow \prod_{t \in I_c} \exp \left\{ m \left[\frac{\epsilon_t}{r_t} \log \left(\frac{er_t A_{r_t}^{(t)} \gamma_t}{\epsilon_t} \right) + O(\epsilon_t^2) \right] \right\} = \exp \left\{ m \left[\sum_{t \in I_c} \left(\frac{\epsilon_t}{r_t} \log \left(\frac{e \rho_t C_t}{\epsilon_t \int \rho} \right) \right) + O(\delta^2) \right] \right\} \quad (15)$$

The number of check-valid assignments of weight δm , which we denote $N_c(\delta m)$, is equal to the sum of $N_c^{(\epsilon)}(\delta m)$ over all admissible vectors ϵ . However, the asymptotic expression as $n \rightarrow \infty$ will be dominated by the distribution ϵ which maximizes the argument of the exponential². Therefore, our next step is to maximize the function

$$f(\epsilon) = \sum_{t \in I_c} \frac{\epsilon_t}{r_t} \log \left(\frac{e \rho_t C_t}{\epsilon_t \int \rho} \right)$$

subject to the constraints

$$g(\epsilon) = \sum_{t \in I_c} \epsilon_t = \delta \quad (16)$$

and

$$\epsilon_t \geq 0 \quad \forall t \in I_c. \quad (17)$$

We solve this optimization problem using Lagrange multipliers, ignoring for the moment the final constraint. Since

$$\frac{\partial f}{\partial \epsilon_t} = \frac{1}{r_t} \log \left(\frac{\rho_t C_t}{\epsilon_t \int \rho} \right) \quad ; \quad \frac{\partial g}{\partial \epsilon_t} = 1$$

for all $t \in I_c$, we have to solve the n_c equations (where μ is the Lagrange multiplier)

$$\frac{1}{r_t} \log \left(\frac{\rho_t C_t}{\epsilon_t \int \rho} \right) + \mu = 0 \quad \forall t \in I_c \quad (18)$$

together with (16), for the $(n_c + 1)$ unknowns $\{\mu, \epsilon\}$. First, (18) yields

$$\epsilon_t = \frac{\rho_t C_t}{\int \rho} z^{r_t} \quad \forall t \in I_c$$

²Observe that as $m \rightarrow \infty$, $\sum_t \exp(m Z_t) \rightarrow \exp(m \max_t \{Z_t\})$

where $z = e^\mu$. Now using (16), we obtain

$$\frac{1}{\int \rho} \sum_{t \in I_c} C_t \rho_t z^{r t} = \delta.$$

The left-hand side of this equation involves a sum of positive terms. For δ sufficiently small, we may approximate

$$\frac{1}{\int \rho} \sum_{t \in I_c} \rho_t C_t z^{r t} \approx \frac{1}{\int \rho} \sum_{t \in X_c} \rho_t C_t z^{r t} \quad (19)$$

$$= \frac{C}{\int \rho} z^r. \quad (20)$$

Applying this approximation, we obtain $\epsilon_t = 0$ if $t \notin X_c$, and $\epsilon_t = K \rho_t C_t$ if $t \in X_c$, where K is independent of t . Then (16) yields $K = \delta/C$, and we obtain the solution

$$\epsilon_t = \begin{cases} \rho_t C_t \delta / C & \text{if } t \in X_c \\ 0 & \text{otherwise.} \end{cases}$$

This solution satisfies (17). When substituted into (15), it yields the following result: as $n \rightarrow \infty$

$$N_c(\delta m) \rightarrow \exp \left\{ m \left[\frac{\delta}{r} \log \left(\frac{eC}{\delta \int \rho} \right) + O(\delta^2) \right] \right\} \quad (21)$$

C. Number of variable-valid split assignments of split weight $(\tau n, \sigma n)$ over γn VNs of type $t \in I_v$

Consider γn VNs of the same type $t \in I_v$. We now evaluate the number of variable-valid split assignments (over these VNs) of split weight $(\tau n, \sigma n)$. Using generating functions³, this is given by

$$N_{v,t}^{(\gamma n)}(\tau n, \sigma n) = \text{Coeff} \left[\left(B^{(t)}(x, y) \right)^{\gamma n}, x^{\tau n} y^{\sigma n} \right]$$

where $\text{Coeff} [p(x, y), x^c y^d]$ denotes the coefficient of $x^c y^d$ in the bivariate polynomial $p(x, y)$. We make use of the following lemma from [10, Theorem 2].

Lemma 4.3: Let

$$B(x, y) = 1 + \sum_{u=1}^k \sum_{v=c}^d B_{u,v} x^u y^v$$

where $k \geq 1$ and $1 \leq c \leq d$, be a bivariate polynomial satisfying $B_{u,v} \geq 0$ for all $1 \leq u \leq k, c \leq v \leq d$. For fixed positive rational numbers ξ and θ , consider the set of positive integers ℓ such that $\xi \ell \in \mathbb{Z}, \theta \ell \in \mathbb{Z}$ and $\text{Coeff} [(B(x, y))^\ell, x^{\xi \ell} y^{\theta \ell}] > 0$. Then either this set is empty, or has infinite cardinality; if t is one such ℓ , then so is $j t$ for every positive integer j . Assuming the latter case, the following limit is well defined and exists:

$$\begin{aligned} \lim_{\ell \rightarrow \infty} \frac{1}{\ell} \log \text{Coeff} \left[(B(x, y))^\ell, x^{\xi \ell} y^{\theta \ell} \right] \\ = \max_{\boldsymbol{\eta}} \sum_{(i,j) \in S} \eta_{i,j} \log \left(\frac{B_{i,j}}{\eta_{i,j}} \right) \end{aligned} \quad (22)$$

³We use the following result on bivariate generating functions [20]. Let $a_{i,j}$ be the number of ways of obtaining an outcome $(i, j) \in \mathbb{Z}^2$ in experiment \mathcal{A} , and let $b_{k,l}$ be the number of ways of obtaining an outcome $(k, l) \in \mathbb{Z}^2$ in experiment \mathcal{B} . Also let $c_{p,q}$ be the number of ways of obtaining an outcome $((i, j), (k, l))$ in the combined experiment $(\mathcal{A}, \mathcal{B})$ with sums $i+k = p$ and $j+l = q$. Then the generating functions $A(x, y) = \sum_{i,j} a_{i,j} x^i y^j, B(x, y) = \sum_{k,l} b_{k,l} x^k y^l$ and $C(x, y) = \sum_{p,q} c_{p,q} x^p y^q$ are related by $C(x, y) = A(x, y)B(x, y)$.

where $S = \{(i, j) \in \mathbb{Z}^2 : B_{i,j} > 0\}, \boldsymbol{\eta} = (\eta_{i,j})_{(i,j) \in S}$, and the maximization is subject to the constraints $\sum_{(i,j) \in S} \eta_{i,j} = 1, \sum_{(i,j) \in S} i \eta_{i,j} = \xi, \sum_{(i,j) \in S} j \eta_{i,j} = \theta$ and $\eta_{i,j} \geq 0$ for all $(i, j) \in S$.

Applying this lemma by substituting $B(x, y) = B^{(t)}(x, y), l = \gamma n, \xi = \tau/\gamma$ and $\theta = \sigma/\gamma$, we obtain that with γ fixed, as $n \rightarrow \infty$

$$N_{v,t}^{(\gamma n)}(\tau n, \sigma n) = \text{Coeff} \left[\left(B^{(t)}(x, y) \right)^{\gamma n}, x^{\tau n} y^{\sigma n} \right] \quad (23)$$

$$\rightarrow \exp \left\{ n \gamma \max_{\boldsymbol{\eta}^{(t)}} \sum_{(i,j) \in S_t} \eta_{i,j}^{(t)} \log \left(\frac{B_{i,j}^{(t)}}{\eta_{i,j}^{(t)}} \right) \right\} \quad (24)$$

$$\triangleq \exp \left\{ n X_t^{(\gamma)}(\tau, \sigma) \right\} \quad (25)$$

where the maximization over $\boldsymbol{\eta}^{(t)} = (\eta_{i,j}^{(t)})_{(i,j) \in S_t}$ is subject to the constraints $\sum_{(i,j) \in S_t} \eta_{i,j}^{(t)} = 1, \sum_{(i,j) \in S_t^-} i \eta_{i,j}^{(t)} = \tau/\gamma, \sum_{(i,j) \in S_t^-} j \eta_{i,j}^{(t)} = \sigma/\gamma$ and $\eta_{i,j}^{(t)} \geq 0$ for all $(i, j) \in S_t$ (recall that the sets S_t and S_t^- are given by (8) and (9)).

D. Growth rate of the weight distribution of the irregular D-GLDPC code ensemble sequence

Recall that the number of check-valid assignments of weight δm is $N_c(\delta m)$; also, the total number of assignments of weight δm is $\binom{E}{\delta m}$. Therefore, the probability that a randomly chosen assignment of weight δm is check-valid is given by

$$P_{\text{valid}}(\delta m) = N_c(\delta m) / \binom{E}{\delta m}.$$

Here we adopt the notation $\delta m = \beta n$; also we have $E = m/\int \rho = n/\int \lambda$. The binomial coefficient may be asymptotically approximated using the fact, based on Stirling's approximation, that as $n \rightarrow \infty$ [11]

$$\binom{\tau n}{\sigma n} \rightarrow \exp \left\{ n \left[\sigma \log \left(\frac{e\tau}{\sigma} \right) + O(\sigma^2) \right] \right\}$$

(valid for $0 < \sigma < \tau < 1$) which yields, in this case,

$$\binom{n/\int \lambda}{\beta n} \rightarrow \exp \left\{ n \left[\beta \log \left(\frac{e}{\beta \int \lambda} \right) + O(\beta^2) \right] \right\}$$

as $n \rightarrow \infty$. Applying this together with the asymptotic expression (21), and assuming sufficiently small β , we find that as $n \rightarrow \infty$ (exploiting the fact that $\delta \int \rho = \beta \int \lambda$)

$$P_{\text{valid}}(\beta n) \rightarrow \exp\{nY(\beta)\} \quad (26)$$

where

$$Y(\beta) = \frac{\beta}{r} \log \left(\frac{eC}{\beta \int \lambda} \right) - \beta \log \left(\frac{e}{\beta \int \lambda} \right) + O(\beta^2).$$

Next, we note that the expected number of D-GLDPC codewords of weight αn in the ensemble \mathcal{M}_n is equal to the sum over β of the expected numbers of split assignments of split weight $(\alpha n, \beta n)$ which are both check-valid and variable-valid, denoted $N_{\alpha n, \beta n}^{v,c}$:

$$\mathbb{E}_{\mathcal{M}_n} [N_{\alpha n}] = \sum_{\beta} \mathbb{E}_{\mathcal{M}_n} [N_{\alpha n, \beta n}^{v,c}].$$

This may then be expressed as

$$\mathbb{E}_{\mathcal{M}_n} [N_{\alpha n}] = \sum_{\beta} P_{\text{valid}}(\beta n) \sum_{\substack{\sum \alpha_t = \alpha \\ \sum \beta_t = \beta}} \left[\prod_{t \in I_v} N_{v,t}^{(\delta_t n)}(\alpha_t n, \beta_t n) \right]$$

where the fraction δ_t of VNs of type $t \in I_v$ is given by (5) and the second sum is over all partitions of α and β into n_v elements, i.e., we have $\alpha_t, \beta_t \geq 0$ for all $t \in I_v$, and $\sum_{t \in I_v} \alpha_t = \alpha$, $\sum_{t \in I_v} \beta_t = \beta$.

Now, using (23)-(25), as $n \rightarrow \infty$ we have for each $t \in I_v$

$$N_{v,t}^{(\delta_t n)}(\alpha_t n, \beta_t n) \rightarrow \exp \left\{ n X_t^{(\delta_t)}(\alpha_t, \beta_t) \right\},$$

where, for each $t \in I_v$,

$$X_t^{(\delta_t)}(\alpha_t, \beta_t) = \delta_t \max_{\boldsymbol{\eta}^{(t)}} \sum_{(i,j) \in S_t} \eta_{i,j}^{(t)} \log \left(\frac{B_{i,j}^{(t)}}{\eta_{i,j}^{(t)}} \right) \quad (27)$$

and the maximization over $\boldsymbol{\eta}^{(t)} = (\eta_{i,j}^{(t)})_{(i,j) \in S_t}$ is subject to the constraints

$$\sum_{(i,j) \in S_t} \eta_{i,j}^{(t)} = 1 \quad (28)$$

$$\sum_{(i,j) \in S_t^-} i \eta_{i,j}^{(t)} = \alpha_t / \delta_t \quad (29)$$

$$\sum_{(i,j) \in S_t^-} j \eta_{i,j}^{(t)} = \beta_t / \delta_t \quad (30)$$

and

$$\eta_{i,j}^{(t)} \geq 0 \quad \forall (i,j) \in S_t. \quad (31)$$

Therefore, recalling (26), we have that as $n \rightarrow \infty$,

$$\mathbb{E}_{\mathcal{M}_n} [N_{\alpha n}] \rightarrow \sum_{\beta} \sum_{\substack{\sum \alpha_t = \alpha \\ \sum \beta_t = \beta}} \exp \left\{ n \left[\sum_{t \in I_v} X_t^{(\delta_t)}(\alpha_t, \beta_t) + Y(\beta) \right] \right\}. \quad (32)$$

Next, for each $t \in I_v$ we define

$$F_t(\boldsymbol{\eta}^{(t)}) = \eta_{0,0}^{(t)} \log \left(\frac{1}{\eta_{0,0}^{(t)}} \right) - \sum_{(i,j) \in S_t^-} \eta_{i,j}^{(t)}.$$

Note that the expression (32) is dominated as $n \rightarrow \infty$ by the term which maximizes the argument of the exponential. Thus we may write

$$G(\alpha) = \max_{\beta} \max_{\substack{\sum \alpha_t = \alpha \\ \sum \beta_t = \beta}} \left\{ \sum_{t \in I_v} \delta_t \max_{\boldsymbol{\eta}^{(t)}} \left[\sum_{(i,j) \in S_t^-} \eta_{i,j}^{(t)} \log \left(\frac{e B_{i,j}^{(t)}}{\eta_{i,j}^{(t)}} \right) + F_t(\boldsymbol{\eta}^{(t)}) \right] + \frac{\beta}{r} \log \left(\frac{eC}{\beta f \lambda} \right) - \beta \log \left(\frac{e}{\beta f \lambda} \right) + O(\beta^2) \right\} \quad (33)$$

where the maximization over $\boldsymbol{\eta}^{(t)} = (\eta_{i,j}^{(t)})_{(i,j) \in S_t^-}$ (for each $t \in I_v$) is subject to constraints (29) and (30) together with $\eta_{i,j}^{(t)} \geq 0$ for all $(i,j) \in S_t^-$.

We next have the following lemma.

Lemma 4.4: The expression $\sum_{t \in I_v} \delta_t F_t(\boldsymbol{\eta}^{(t)})$ is $O(\alpha^2)$ for any $\boldsymbol{\eta}^{(t)}$ satisfying the optimization constraints (28)-(31)⁴.

The proof of this lemma is omitted due to space considerations. It follows from Lemma 4.4 that the expression $\sum_{t \in I_v} \delta_t F_t(\boldsymbol{\eta}^{(t)})$ is $O(\alpha^2)$ for the maximizing $\boldsymbol{\eta}^{(t)}$. Also, since β/α is bounded between two positive constants, any expression which is $O(\beta^2)$ must necessarily also be $O(\alpha^2)$. Therefore

$$G(\alpha) = \max_{\beta} \max_{\substack{\sum \alpha_t = \alpha \\ \sum \beta_t = \beta}} \left[\sum_{t \in I_v} \delta_t \max_{\boldsymbol{\eta}^{(t)}} \sum_{(i,j) \in S_t^-} \eta_{i,j}^{(t)} \log \left(\frac{e B_{i,j}^{(t)}}{\eta_{i,j}^{(t)}} \right) + \frac{\beta}{r} \log \left(\frac{eC}{\beta f \lambda} \right) - \beta \log \left(\frac{e}{\beta f \lambda} \right) \right] + O(\alpha^2)$$

where the optimization is (as before) subject to the constraints (29) and (30) together with $\eta_{i,j}^{(t)} \geq 0$ for all $(i,j) \in S_t^-$. In what follows, for convenience of presentation we shall temporarily omit the $O(\alpha^2)$ term in the expression for growth rate.

Next we make the substitution $\gamma_{i,j}^{(t)} = \delta_t \eta_{i,j}^{(t)}$ for all $t \in I_v$, $(i,j) \in S_t^-$. This yields

$$G(\alpha) = \max_{\beta} \max_{\substack{\sum \alpha_t = \alpha \\ \sum \beta_t = \beta}} \left[\sum_{t \in I_v} \max_{\boldsymbol{\gamma}^{(t)}} \sum_{(i,j) \in S_t^-} \gamma_{i,j}^{(t)} \log \left(\frac{e B_{i,j}^{(t)} \delta_t}{\gamma_{i,j}^{(t)}} \right) + \frac{\beta}{r} \log \left(\frac{eC}{\beta f \lambda} \right) - \beta \log \left(\frac{e}{\beta f \lambda} \right) \right]$$

where the maximization over $\boldsymbol{\gamma}^{(t)} = (\gamma_{i,j}^{(t)})_{(i,j) \in S_t^-}$ (for each $t \in I_v$) is subject to the constraints $\sum_{(i,j) \in S_t^-} i \gamma_{i,j}^{(t)} = \alpha_t$, $\sum_{(i,j) \in S_t^-} j \gamma_{i,j}^{(t)} = \beta_t$, and $\gamma_{i,j}^{(t)} \geq 0$ for all $(i,j) \in S_t^-$. We observe that this maximization may be recast as

$$G(\alpha) = \max_{\boldsymbol{\gamma}} \left[\sum_{t \in I_v} \sum_{(i,j) \in S_t^-} \gamma_{i,j}^{(t)} \log \left(\frac{e B_{i,j}^{(t)} \delta_t}{\gamma_{i,j}^{(t)}} \right) + \frac{\beta(\boldsymbol{\gamma})}{r} \log \left(\frac{eC}{\beta(\boldsymbol{\gamma}) f \lambda} \right) - \beta(\boldsymbol{\gamma}) \log \left(\frac{e}{\beta(\boldsymbol{\gamma}) f \lambda} \right) \right]$$

where the maximization, which is now over $\boldsymbol{\gamma} = (\gamma_{i,j}^{(t)})_{t \in I_v, (i,j) \in S_t^-}$, is subject to the constraints

$$\sum_{t \in I_v} \sum_{(i,j) \in S_t^-} i \gamma_{i,j}^{(t)} = \alpha$$

⁴Here we use the following standard notation: the real-valued function $f(x)$ is said to be $O(g(x))$ if and only if there exist positive real numbers k and ϵ , both independent of x , such that

$$|f(x)| \leq k g(x) \quad \forall 0 \leq x \leq \epsilon.$$

and $\gamma_{i,j}^{(t)} \geq 0$ for all $t \in I_v$, $(i,j) \in S_t^-$, and where

$$\beta(\gamma) = \sum_{t \in I_v} \sum_{(i,j) \in S_t^-} j \gamma_{i,j}^{(t)}.$$

Making the substitution $\nu_{i,j}^{(t)} = \gamma_{i,j}^{(t)}/\alpha$ for all $t \in I_v$, $(i,j) \in S_t^-$, we obtain

$$G(\alpha) = \alpha \max_{\nu} \left[\sum_{t \in I_v} \sum_{(i,j) \in S_t^-} \nu_{i,j}^{(t)} \log \left(\frac{e B_{i,j}^{(t)} \delta_t}{\alpha \nu_{i,j}^{(t)}} \right) + \frac{z(\nu)}{r} \log \left(\frac{eC}{\alpha z(\nu) \int \lambda} \right) - z(\nu) \log \left(\frac{e}{\alpha z(\nu) \int \lambda} \right) \right]$$

where the maximization over $\nu = (\nu_{i,j}^{(t)})_{t \in I_v, (i,j) \in S_t^-}$ is subject to the constraints $\sum_{t \in I_v} \sum_{(i,j) \in S_t^-} i \nu_{i,j}^{(t)} = 1$ and $\nu_{i,j}^{(t)} \geq 0$ for all $t \in I_v$, $(i,j) \in S_t^-$, and where

$$z(\nu) = \sum_{t \in I_v} \sum_{(i,j) \in S_t^-} j \nu_{i,j}^{(t)}.$$

Assuming the condition $r = 2$, we may write

$$G(\alpha) = \alpha \max_{\nu} (K_1(\nu) + K_2(\nu) \log \alpha)$$

where

$$K_2(\nu) = \frac{z(\nu)}{2} - \sum_{t \in I_v} \sum_{(i,j) \in S_t^-} \nu_{i,j}^{(t)}.$$

Next, assuming the condition $p = 2$, we make the observation that

$$z(\nu) = \sum_{t \in I_v} \sum_{(i,j) \in S_t^-} j \nu_{i,j}^{(t)} \geq 2 \sum_{t \in I_v} \sum_{(i,j) \in S_t^-} \nu_{i,j}^{(t)}$$

with equality if and only if $\nu_{i,j}^{(t)} = 0$ for all $t \in I_v$, $(i,j) \in S_t^-$ with $j > 2$. Therefore $K_2(\nu) \geq 0$ with equality if and only if $\nu_{i,j}^{(t)} = 0$ for all $t \in I_v$, $(i,j) \in S_t^-$ with $j > 2$. Let ν_1 and ν_2 be two distributions satisfying the optimization constraints, and suppose that $K_2(\nu_1) > 0$ and $K_2(\nu_2) = 0$. Then for sufficiently small α , we must have

$$K_1(\nu_1) + K_2(\nu_1) \log \alpha < K_1(\nu_2) + K_2(\nu_2) \log \alpha.$$

This follows from the fact that the inequality

$$K_2(\nu_1) \log \alpha < K_1(\nu_2) - K_1(\nu_1)$$

will always be satisfied for α sufficiently small (since $\log \alpha \rightarrow -\infty$ as $\alpha \rightarrow 0$, and recalling that $K_2(\nu_1) > 0$).

Therefore, for sufficiently small α , the vector ν which maximizes $K_1(\nu) + K_2(\nu) \log \alpha$ must satisfy $\nu_{i,j}^{(t)} = 0$ for all $t \in I_v$, $(i,j) \in S_t^-$ with $j > 2$. Note that this implies that the maximum, and hence the growth rate, depends only on the check and VNs with minimum distance equal to 2. Also, recall that for each $t \in X_v$, the set $L_t = \{i \in \mathbb{Z} : B_{i,2}^{(t)} > 0\}$; we contract the vector ν to include only variables $\nu_{i,2}^{(t)}$ where $t \in X_v$ and $i \in L_t$ (since only these may assume positive values).

The growth rate may be written as

$$G(\alpha) = \alpha \max_{\nu} \left[\sum_{t \in X_v} \sum_{i \in L_t} \nu_{i,2}^{(t)} \log \left(\frac{B_{i,2}^{(t)} \delta_t}{\nu_{i,2}^{(t)}} \right) + s(\nu) \log \frac{s(\nu)}{\phi} \right] \triangleq \alpha \max_{\nu} (\log R(\nu)) \quad (34)$$

where

$$\phi \triangleq \frac{1}{2C \int \lambda} \quad (35)$$

and where the function $s(\nu)$ is given by

$$s(\nu) = \sum_{t \in X_v} \sum_{i \in L_t} \nu_{i,2}^{(t)}. \quad (36)$$

The maximization over $\nu = (\nu_{i,2}^{(t)})_{t \in X_v, i \in L_t}$ in (34) is subject to the constraints

$$h(\nu) = \sum_{t \in X_v} \sum_{i \in L_t} i \nu_{i,2}^{(t)} = 1 \quad (37)$$

and

$$\nu_{i,2}^{(t)} \geq 0 \quad \forall t \in X_v, i \in L_t. \quad (38)$$

Let the vector ν which maximizes (34) be denoted by $\tilde{\nu}$. Then, our task is to show that

$$R(\tilde{\nu}) \triangleq \tilde{R} = \frac{1}{P^{-1}(1/C)}$$

i.e., that

$$P\left(\frac{1}{\tilde{R}}\right) = \frac{1}{C}, \quad (39)$$

where the parameter C and the polynomial $P(x)$ are defined in (10) and (11), respectively. We show this using Lagrange multipliers, ignoring for the moment the constraint (38). We have

$$\frac{\partial \log R(\nu)}{\partial \nu_{i,2}^{(t)}} = \log \left(\frac{B_{i,2}^{(t)} \delta_t}{\nu_{i,2}^{(t)}} \right) + \log \left(\frac{s(\nu)}{\phi} \right); \quad \frac{\partial h(\nu)}{\partial \nu_{i,2}^{(t)}} = i$$

so that, at the maximum,

$$\log \left(\frac{B_{i,2}^{(t)} \delta_t}{\tilde{\nu}_{i,2}^{(t)}} \right) + \log \left(\frac{s(\tilde{\nu})}{\phi} \right) = \lambda i$$

for all $t \in X_v$, $i \in L_t$, and for some Lagrange multiplier $\lambda \in \mathbb{R}$. Substituting back into (34) and using (37) yields

$$\log \tilde{R} = \sum_{t \in X_v} \sum_{i \in L_t} \nu_{i,2}^{(t)} \left(\lambda i - \log \left(\frac{s(\tilde{\nu})}{\phi} \right) \right) + s(\tilde{\nu}) \log \left(\frac{s(\tilde{\nu})}{\phi} \right) = \lambda \quad (40)$$

i.e., the maximum value of the function $\log R(\nu)$ is equal to the Lagrange multiplier. Thus we have

$$\frac{B_{i,2}^{(t)}}{(\tilde{R})^i} = \left(\frac{\phi}{s(\tilde{\nu}) \delta_t} \right) \tilde{\nu}_{i,2}^{(t)}$$

for all $t \in X_v, i \in L_t$. Substituting this into the LHS of (39) and recalling the definition (11), we obtain

$$\begin{aligned} P\left(\frac{1}{\tilde{R}}\right) &= \sum_{t \in X_v} \sum_{i \in L_t} \frac{2\lambda_t B_{i,2}^{(t)}}{qt (\tilde{R})^i} \\ &= \sum_{t \in X_v} \sum_{i \in L_t} \frac{\tilde{\nu}_{i,j}^{(t)}}{Cs(\tilde{\nu})} = \frac{1}{C} \end{aligned}$$

where we have used (5), (35) and (36). This completes the proof of the theorem. Note that (12) is a first-order Taylor series expansion around $\alpha = 0$ which directly generalizes the results of [11] and [15] (for irregular LDPC and GLDPC codes respectively) to the case of irregular D-GLDPC codes. Our result indicates that for this case also, the parameter $1/P^{-1}(1/C)$ plays an analogous role to the parameter $\lambda'(0)\rho'(1)$ for irregular LDPC codes, and to the parameter $\lambda'(0)C$ for irregular GLDPC codes.

V. CONCLUSION

An expression for the asymptotic growth rate of the weight distribution of D-GLDPC codes for small linear-weight codewords has been derived. The expression assumes the existence of minimum distance 2 check and variable nodes, and involves the evaluation of a polynomial inverse, derived from the minimum distance 2 variable nodes, at a point derived from the minimum distance 2 check nodes. This generalizes known results for LDPC codes and GLDPC codes, and also generalizes the corresponding connection with the stability condition over the BEC.

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